Advanced Approximate Inference for Discrete LVMs

Deep Learning 2 – 2022

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2 Sparse Gradients

3 Variance reduction

# Gradients of the evidence lowerbound

#### ELBO

$$\mathcal{E}(\lambda, heta) = \mathbb{E}_{q_{Z|X}(z|x,\lambda)} \left[ \log rac{p_{ZX}(z,x| heta)}{q_{Z|X}(z|x,\lambda)} 
ight]$$

Updating generative model

$$\boldsymbol{\nabla}_{\theta} \mathcal{E}(\lambda, \theta) = \mathbb{E}_{\boldsymbol{q}_{Z|X}(\boldsymbol{z}|\boldsymbol{x}, \lambda)} \left[ \boldsymbol{\nabla}_{\theta} \log \frac{\boldsymbol{p}_{ZX}(\boldsymbol{z}, \boldsymbol{x}|\theta)}{\boldsymbol{q}_{Z|X}(\boldsymbol{z}|\boldsymbol{x}, \lambda)} \right]$$

Updating inference model

 $\boldsymbol{\nabla}_{\lambda} \mathcal{E}(\lambda, \theta) = \boldsymbol{\nabla}_{\lambda} \mathbb{E}_{f_{Z}(\boldsymbol{z}|\lambda)} \left[ \psi(\boldsymbol{z}) \right]$ 

$$f_Z(z|\lambda) = q_{Z|X}(z|x,\lambda)$$
 and  $\psi(z) = \log rac{p_{ZX}(z,x| heta)}{q_{Z|X}(z|x,\lambda)}$ 

# Score function estimator

$$\frac{\partial}{\partial \lambda} \mathbb{E}_{f_{Z|\lambda}(z)} \left[ \psi(z) \right] = \mathbb{E}_{f_{Z|\lambda}(z)} \left[ \psi(z) \frac{\partial}{\partial \lambda} \log f_{Z|\lambda}(z) \right]$$

Easy to MC estimate, but noisy.

# Reparameterised gradient

$$\frac{\partial}{\partial\lambda} \mathbb{E}_{f_{Z|\lambda}(z)} \left[ \psi(z) \right] = \mathbb{E}_{s(\epsilon)} \left[ \frac{\partial}{\partial\lambda} \psi(t(\epsilon, \lambda)) \right]$$
$$= \mathbb{E}_{s(\epsilon)} \left[ \frac{\partial}{\partial z} \psi(z) \frac{\partial}{\partial\lambda} t(\epsilon, \lambda) \right]$$

Easy to MC estimate.

#### Reparameterisation

- $t(\epsilon, \lambda)$  is invertible and differentiable
- $z = t(\epsilon, \lambda)$  has density  $f_Z(z|\lambda)$
- $\epsilon = t^{-1}(z, \lambda)$  has density  $s(\epsilon)$

Change of density

$$f_{Z|\lambda}(z) = s(\underbrace{t^{-1}(z,\lambda)}_{\epsilon}) |\det J_{t^{-1}}(z,\lambda)|$$

As a result

$$\begin{split} \mathbb{E}_{f_{Z}(z|\lambda)}[\psi(z)] &= \int f_{Z}(z|\lambda)\psi(z)\mathrm{d}z \\ &= \int s(t^{-1}(z,\lambda))|\mathrm{det}\,J_{t^{-1}}(z,\lambda)|\psi(z)\mathrm{d}z \\ &= \int s(\epsilon)|\mathrm{det}\,J_{t}(\epsilon,\lambda)|^{-1}\psi(t(\epsilon,\lambda))|\mathrm{det}\,J_{t}(\epsilon,\lambda)|\mathrm{d}\epsilon \\ &= \int s(\epsilon)\psi(t(\epsilon,\lambda))\mathrm{d}\epsilon \end{split}$$

### Gradient Estimators

Basic problem: we want to differentiate an expected value wrt  $\lambda$ 

$$\frac{\partial}{\partial \lambda} \mathbb{E}_{f_{Z|\lambda}}[\psi(z)] \qquad \text{e.g., } \psi(z) := \log p(x|z,\theta)$$
$$f_{Z|\lambda}(z) := q(z|x,\lambda)$$

but the distribution of Z depends on  $\lambda$ .

We have met the SFE and the reparameterised gradient estimator:

$$\mathbb{E}_{f_{\lambda}(z)}\left[\underbrace{\psi(z)\frac{\partial}{\partial\lambda}\log f_{Z|\lambda}(z)}_{\hat{g}_{sfe}}\right] = \mathbb{E}_{s(\epsilon)}\left[\underbrace{\frac{\partial}{\partial z}\psi(z)\frac{\partial}{\partial\lambda}t(\epsilon,\lambda)}_{\hat{g}_{rep}}\right]$$

•  $\hat{g}_{sfe}$  is typically cursed with variance

From VAEs, you know the reparameterised gradient estimator

$$\begin{split} \frac{\partial}{\partial \lambda} \mathbb{E}_{f_{Z|\lambda}(z)} \left[ \psi(z) \right] &= \mathbb{E}_{s(\epsilon)} \left[ \frac{\partial}{\partial \lambda} \psi(t(\epsilon, \lambda)) \right] \\ &= \mathbb{E}_{s(\epsilon)} \left[ \frac{\partial}{\partial z} \psi(z) \frac{\partial}{\partial \lambda} t(\epsilon, \lambda) \right] \end{split}$$

it takes an invertible and differentiable transformation t such that

$$t(\epsilon,\lambda) \sim f_{Z|\lambda} \ t^{-1}(z,\lambda) \sim s(\epsilon)$$

**Goals** Understand why there can't be a  $\hat{g}_{rep}$  for discrete rvs. Meet alternatives to SFE.

# A general reparameterisation

For univariate Z, what transformation will always absorb the parameters of the density  $f_{Z|\lambda}(z)$ ?

The argument extends to a vector of independent univariate rvs.

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So, if I know the inverse cdf,

$$\epsilon \sim \mathcal{U}(0,1)$$
 $F_{Z|\lambda}^{-1}(\epsilon) \sim Z|\lambda$ 

I have access to  $\hat{g}_{rep}$ .

The argument extends to a vector of independent univariate rvs.



 $Z \sim \text{Bernoulli}(p)$ 



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How about  $\frac{\partial}{\partial p} F_{Z|p}^{-1}(\epsilon)$ ?



How about  $\frac{\partial}{\partial p} F_{Z|p}^{-1}(\epsilon)$ ? Mostly 0, sometimes undefined!

#### Discrete case

Discrete variables do not admit a differentiable reparameterisation. The derivatives of the inverse cdf are either 0 or undefined :/

STE's original paper (Bengio et al., 2013).

There are other pseudo-gradients in the literature, for example for relaxed combinatorial random variables (Peng et al., 2018; Mihaylova et al., 2021).

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The score function estimator is fully general, but very noisy.

How about we fake a Jacobian and call it a pseudo-gradient?

 $J_t(\epsilon, \lambda) = \operatorname{diag}(\mathbf{1})$ 

This is the ingredient behind the straight-through estimator (STE).

STE's original paper (Bengio et al., 2013).

There are other pseudo-gradients in the literature, for example for relaxed combinatorial random variables (Peng et al., 2018; Mihaylova et al., 2021).

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Optimising the ELBO via reparameterised samples requires

$$\hat{g}_{\mathsf{rep}} = \frac{\partial}{\partial \lambda} \log p(x|z = t(\epsilon, \lambda)) = \frac{\partial}{\partial z} \log p(x|z) \frac{\partial}{\partial \lambda} t(\epsilon, \lambda)$$

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Let's use our *pseudo gradient* 

$$\hat{g}_{\mathsf{ste}} := \frac{\partial}{\partial \lambda} t(\epsilon, \lambda) = \frac{\partial}{\partial \lambda} g(x; \lambda) \frac{\partial}{\partial p} \mathbb{1}_{(0,p)}(\epsilon)$$

-

We can sample from a Categorical distribution via

Concrete distribution (Maddison et al., 2017), Gumbel-Softmax distribution (Jang et al., 2017).

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$$\underbrace{ \substack{ \epsilon_k \sim \mathsf{Gumbel}(0,1) \\ \underset{k}{\operatorname{arg\,max}} \{\lambda_k + \epsilon_k\}_{k=1}^K \sim \mathsf{Cat}(\mathsf{softmax}(\lambda)) \\ \underbrace{ \sum_{k=t(\epsilon,\lambda)}}_{z=t(\epsilon,\lambda)} }$$

The problem is that  $t(\epsilon, \lambda)$  is not differentiable, but note

$$\operatorname{softmax}\left(rac{\lambda+\epsilon}{ au}
ight) o \operatorname{onehot}(z) \qquad \operatorname{as}\, au o \mathsf{0}$$

and now the transformation is differentiable, but the outcome is dense.

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Continuous random variables that take on sparse outcomes with non-zero probability mass.

Example I:

- sample  $\zeta \sim \mathcal{N}(0, 1)$
- rectify the sample via hardsigmoid  $z = \min(1, \max(0, z))$ .

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What is the probability  $\Pr(Z \in \{0\})$ ? it is  $\Phi(0) = \int_{\mathbb{R}_{<0}} \mathcal{N}(0,1) \mathrm{d}z$ 

When  $\zeta < 0$  or  $\zeta > 1$  the derivative of hardsigmoid is 0, when  $0 < \zeta < 1$  the derivative is 1. Hardsigmoid has undefined derivatives for  $\zeta = 0$  and  $\zeta = 1$ , but we will never sample those.

If we had a parameterised Gaussian, we could sample with a reparameterisation and learn the Gaussian parameters.

This has been applied to generate mixed random variables in the support [0, 1].

Spike-and-slab (Rolfe, 2017); HardConcrete (Louizos et al., 2018); Hard-Kumaraswamy (Bastings et al., 2019).

Applications to interpretability (Voita et al., 2019; Cao et al., 2020; Ataman et al., 2020)

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Continuous random variables that take on sparse outcomes with non-zero probability mass.

Construction:

- start with a continuous univariate rv  $\zeta$  with density  $s(\zeta)$
- $\bullet$  project it to [0,1] using a function that hits the boundaries of the set
- the projection function is differentiable everywhere except at ζ = 0 and ζ = 1 which have 0 measure under s(ζ)

The path derivative is defined *almost everywhere*, a g-rep can be used.

### Can we go beyond univariates?

Yes, if we look into sparse projections to the probability simplex.

It turns out the 'hard sigmoid' is the a special case of a more general projection known as *sparsemax*.

Sparsemax





When K = 2, sparsemax is equivalent to hardsigmoid.

## Mixed Random Variables

Generalisation of mixed binary random variables to the multivariate case.

Intrinsic view:

- draw  $\zeta$  in  $\mathbb{R}^{K}$  (e.g., from a multivariate Gaussian)
- project it to  $\Delta_{K-1}$  (e.g., using sparsemax)

What is the probability that we have a point in one of the faces of the simplex?

The faces of the simplex (e.g., with 3 vertices):



## Mixed Random Variables

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Intrinsic view:

- draw  $\zeta$  in  $\mathbb{R}^{K}$  (e.g., from a multivariate Gaussian)
- project it to  $\Delta_{K-1}$  (e.g., using sparsemax)

What is the probability that we have a point in one of the faces of the simplex? In this example, we would have to integrate the multivariate Gaussian pdf over the set of points in the inverse of sparsemap.

The faces of the simplex (e.g., with 3 vertices):



# Mixed Random Variables

Generalisation of mixed binary random variables to the multivariate case.

Extrinsic view:

- choose a face f of the simplex with probability  $P_F(f)$
- draw *z* from the relative interior of *f* (e.g., using a LogisticNormal distribution)

The faces of the simplex (e.g., with 3 vertices):





Farinhas et al. (2022)



Reparameterised Gradients





#### Latent Computation Graphs

Estimators built on reparameterisation require

- z to be of some fixed finite dimensionality
- the decoder's computation graph must be independent of z.

Some composition functions are parameterised by their inputs (e.g., a tree-LSTM), they are dynamic computation graphs controlled by the discrete latent.

STE is not an option, so we are back to SFE. Or are we?

 $\hat{g}_{rep}$  differentiates the decoder wrt z

$$\hat{g}_{\mathsf{rep}} = rac{\partial}{\partial z} \psi(z) imes rac{\partial}{\partial \lambda} t(\epsilon, \lambda)$$

This cannot work when the computation graph of  $\psi$  depends on z (i.e., whenever z cannot be treated as a point in the relative interior of a fixed and finite-dimensional polytope). For example, a tree-LSTM updates its states following a depth-first traversal of an input tree.

## Parameterise for Tractability

We can use sparse projections to the probability simplex to parameterise discrete distributions that assign 0 probability mass to most of the outcomes in their supports.

$$\begin{aligned} \nabla_{\lambda} \mathbb{E}_{q(z|x,\lambda)} [\log p(x,z|\theta) - \log q(z|x,\lambda)] \\ = & \nabla_{\lambda} q(c_{1}|x,\lambda) (\log p(x,c_{1}|\theta) - \log q(c_{1}|x,\lambda)) \\ & + \dots \\ & + \nabla_{\lambda} q(c_{k}|x,\lambda) (\log p(x,c_{k}|\theta) - \log q(c_{k}|x,\lambda)) \\ & + \nabla_{\lambda} \sum_{z=c_{k+1}}^{c_{k}} \underbrace{q(z|x,\lambda)}_{=0} (\log p(x,z|\theta) - \log q(z|x,\lambda)) \end{aligned}$$

Sparse projections: (Martins and Astudillo, 2016; Niculae et al., 2018a) Latent dynamic computation graphs: (Niculae et al., 2018b) Sparse marginals: (Correia et al., 2020)

Effectively, we have an inference network that parameterises a model whose support is small enough for *enumeration*.

#### Deep Learning 2 @ UvA

#### Advanced discrete LVMs



Reparameterised Gradients

2 Sparse Gradients


#### Control variates

#### Intuition

To estimate  $\mathbb{E}[\psi(z)]$  via Monte Carlo we compute the empirical average of  $\hat{\psi}(z)$  where  $\hat{\psi}(z)$  is chosen so that  $\mathbb{E}[\hat{\psi}(z)] = \mathbb{E}[\psi(z)]$  and  $\operatorname{Var}(\psi) > \operatorname{Var}(\hat{\psi})$ .

Let  $ar{\psi} = \mathbb{E}[\psi(z)]$  be an expectation of interest

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  - and  $\operatorname{Var}(\hat{\psi}) = \operatorname{Var}(\psi) 2b\operatorname{Cov}(\psi, c) + b^2\operatorname{Var}(c)$

# Choosing the control variate

• 
$$\hat{\psi}(z) \triangleq \psi(z) - b(c(z) - \mathbb{E}[c(z)])$$
  
•  $\operatorname{Var}(\hat{\psi}) = \operatorname{Var}(\psi) - 2b \operatorname{Cov}(\psi, c) + b^2 \operatorname{Var}(c)$   
How do we choose *b* and  $c(z)$ ?

#### Choosing the control variate

- $\hat{\psi}(z) \triangleq \psi(z) b(c(z) \mathbb{E}[c(z)])$
- **2**  $\operatorname{Var}(\hat{\psi}) = \operatorname{Var}(\psi) 2b \operatorname{Cov}(\psi, c) + b^2 \operatorname{Var}(c)$
- How do we choose b and c(z)?
  - solving  $\frac{\partial}{\partial b} \operatorname{Var}(\hat{\psi}) = 0$  yields  $b^* = \operatorname{Cov}(\psi, c) / \operatorname{Var}(c)$ when  $\psi(z)$  and c(z) are positively correlated, then we may reduce variance

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- Of course,  $\mathbb{E}[c(z)]$  must be known!

# MC

We then use the estimate

$$ar{\psi} \stackrel{\mathsf{MC}}{\approx} rac{1}{\mathcal{S}} \left( \sum_{s=1}^{\mathcal{S}} \psi(z^{(s)}) - bc(z^{(s)}) \right) + bar{c}$$

## MC

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$$\bar{\psi} \stackrel{\mathsf{MC}}{\approx} \frac{1}{S} \left( \sum_{s=1}^{S} \psi(z^{(s)}) - bc(z^{(s)}) \right) + b\bar{c}$$

And recall that for us

$$\psi(z) = \log \frac{p_{ZX}(z, x|\theta)}{q_{Z|X}(z|x, \lambda)} \frac{\partial}{\partial \lambda} \log q_{Z|X}(z|x, \lambda)$$

and  $z^{(s)} \sim q_{Z|X}(z|x,\lambda)$ 

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$$\mathbb{E}_{q(z|x,\lambda)}\left[rac{\partial}{\partial\lambda}\log q(z|x,\lambda)
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Variance re	duction
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## Baselines

With

$$\psi(z) = \log rac{p(z,x| heta)}{q(z|x,\lambda)} rac{\partial}{\partial\lambda} \log q(z|x,\lambda)$$

and

$$c(z) = rac{\partial}{\partial \lambda} \log q(z|x,\lambda)$$

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*b* is known as *baseline* in RL literature.

• Moving average of log  $\frac{p(z,x|\theta)}{q(z|x,\lambda)}$  based on previous batches

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- The reward assessed at a deterministic point, e.g.  $b(x) = \log \frac{p(z^*, x|\theta)}{q(z^*|x, \lambda)} \text{ where } z^* = \arg \max_z q(z|x, \lambda)$

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- The reward assessed at a stochastic point, e.g.  $b(x) = \log \frac{p(z',x|\theta)}{q(z'|x,\lambda)}$  where  $z' \sim q(z|x,\lambda)$

## Trainable baselines

Baselines are predicted by a regression model (e.g. a neural net).

The model is trained using an  $L_2$ -loss.

$$\min_{\omega} \left( b(x;\omega) - \log \frac{p(z,x|\theta)}{q(z|x,\lambda)} \right)^2$$

## Other techniques

- control variates beyond baselines: Tucker et al. (2017), Grathwohl et al. (2018)
- Rao-Blackwell: Liu et al. (2019)

### Summary

Learning discrete LVMs poses challenges for gradient estimation, in particular, gradients of the inference network are challenging.

SFE is the most general, it requires tractable pmf and sampling, nothing else. It is too noisy to be useful without variance reduction techniques.

#### Summary

Alternatives to SFE are possible in some cases.

STE requires a relaxation of the decoder and introduces biases, violating the requirements for stochastic optimisation.

We can mix a pmf and a pdf to obtain reparameterised and unbiased gradients for a sparse rv. This addresses STE's bias.

Sparse parameterisation of the inference model leads to sparse gradients with many terms evaluating trivially to zero. Enumeration dispenses with relaxations and works for combinatorial variables.

### **Final Remarks**

- Probabilistic models are extremely flexible tools.
- They are interesting precisely because we can make choices about unobserved aspects of the data.
- Discrete latent variables are oftentimes key to revealing interpretable structure, or to imposing some interpretable structure on a joint distribution.
- Learning discrete LVMs is challenging, but recent years have seen amazing progress.
- Join the party! Apply these models, extend them, discover problems with their estimation/evaluation, investigate solutions.
- Avoid approaching LVMs wondering whether they will beat some non-LVM NN. If such NN exists, then you are probably looking at an aspect of the problem that does not require latent variables.

### What Next?

- For more material, check https://vitutorial.github.io/classes/
- We also have some coding exercises https://github.com/vitutorial/exercises
- Check this great tutorial by our friends from DeepSPIN https://deep-spin.github.io/tutorial/

See you around!

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